

Agenda:

1. Estimating $\text{Avar}[\hat{\beta}_n]$

2. Testing linear restrictions on β_*

Recall: $\sqrt{n}(\hat{\beta}_n - \beta_*) \xrightarrow{d} N(0, \underbrace{\Sigma_x^{-1} \Sigma_g \Sigma_x^{-1}}_{\text{Avar}[\hat{\beta}_n]})$ | Interpretation: $\hat{\beta}_n \stackrel{d}{\approx} N(\beta_*, \frac{1}{n} \text{Avar}[\hat{\beta}_n])$

$$\Sigma_x = E[x_i x_i^T]$$

$$\Sigma_g = E[\underbrace{(x_i \varepsilon_i)}_{j_i} (x_i \varepsilon_i)^T] = E[\underbrace{\varepsilon_i^2}_{1 \times 1} \underbrace{x_i x_i^T}_{p \times p}] = E[\underbrace{x_i}_{p \times 1} \underbrace{\varepsilon_i^2}_{1 \times 1} \underbrace{x_i^T}_{1 \times p}]$$

Assume consistent estimates of Σ_x & Σ_g : $\hat{\Sigma}_x \xrightarrow{p} \Sigma_x$

$$\hat{\Sigma}_g \xrightarrow{p} \Sigma_g$$

$$\text{Define } \widehat{\text{Avar}}[\hat{\beta}_n] \equiv \hat{\Sigma}_x^{-1} \hat{\Sigma}_g \hat{\Sigma}_x^{-1}$$

$\widehat{\text{Avar}}[\hat{\beta}_n]$ is a consistent estimate of $\text{Avar}[\hat{\beta}_n]$ (consequence of CMT)

Think of $\text{Avar}[\hat{\beta}_n]$ as $f(\Sigma_x, \Sigma_g)$, where $f(A, B) \equiv A^{-1} B A^{-1}$

Assume $\hat{\Sigma}_x, \hat{\Sigma}_g$ are consistent, then we wish to deduce

$$f(\hat{\Sigma}_x, \hat{\Sigma}_g) \rightarrow f(\Sigma_x, \Sigma_g)$$

consistent estimator of Σ_x is $\hat{\Sigma}_x \equiv \frac{1}{n} \sum_{i=1}^n x_i x_i^T$ (consequence of LLN)

consistent estimator of Σ_g is $\hat{\Sigma}_g \equiv \frac{1}{n} \sum_{i=1}^n \hat{\varepsilon}_i^2 x_i x_i^T$, where

$\hat{\varepsilon}_i$'s are the OLS residuals

$$\hat{\varepsilon}_i \equiv y_i - \hat{y}_i = y_i - x_i^T \hat{\beta}_n$$

Intuition: $\hat{\varepsilon}_i \approx \varepsilon_i$ so $\hat{\Sigma}_g \approx \frac{1}{n} \sum_{i=1}^n \varepsilon_i^2 x_i x_i^T \rightarrow \Sigma_g$ (consequence of LLN)

$$\text{Recall: } \hat{\varepsilon}_i = y_i - x_i^T \hat{\beta}_n = x_i^T \beta_* + \varepsilon_i - x_i^T \hat{\beta}_n = \varepsilon_i + x_i^T (\beta_* - \hat{\beta}_n)$$

$$\hat{\Sigma}_g = \frac{1}{n} \sum_{i=1}^n \hat{\varepsilon}_i^2 x_i x_i^T$$

$$= \frac{1}{n} \sum_{i=1}^n (\varepsilon_i + x_i^T (\beta_* - \hat{\beta}_n))^2 x_i x_i^T \quad \left[\begin{array}{l} \text{plugging in} \\ \text{for } \hat{\varepsilon}_i \end{array} \right. \text{exp.}]$$

$$\rightarrow \frac{1}{n} \sum_{i=1}^n \varepsilon_i^2 x_i x_i^T + \frac{2}{n} \sum_{i=1}^n x_i (\varepsilon_i x_i^T (\beta_* - \hat{\beta}_n)) x_i^T$$

$$+ \frac{1}{n} \sum_{i=1}^n x_i x_i^T (\beta_* - \hat{\beta}_n) (\beta_* - \hat{\beta}_n)^T x_i x_i^T$$

$$= \frac{1}{n} \sum_{i=1}^n \varepsilon_i^2 x_i x_i^T + \frac{1}{n} \sum_{i=1}^n x_i^T x_i (\varepsilon_i x_i^T) (\beta_* - \hat{\beta}_n) \quad \text{(trace trick)}$$

$$+ (\hat{\beta}_n - \beta_0)^T \underbrace{\left(\frac{1}{n} \sum_{i=1}^n x_i x_i^T \right)}_{\mathbb{M}} (\beta_0 - \hat{\beta}_n)$$

$I \xrightarrow{p} Z_g$ (by LLN)

All other terms will converge ^{in probability} to 0 because $\hat{\beta}_n - \beta_0 \xrightarrow{p} 0$

In summary, $\widehat{\text{Avar}}[\hat{\beta}_n] \equiv \underbrace{\left(\frac{1}{n} \sum_{i=1}^n x_i x_i^T \right)^{-1}}_{\hat{\Sigma}_x} \underbrace{\left(\frac{1}{n} \sum_{i=1}^n \hat{e}_i^2 x_i x_i^T \right)}_{\hat{\Sigma}_g} \left(\frac{1}{n} \sum_{i=1}^n x_i x_i^T \right)^{-1}$

is a consistent estimator of $\text{Avar}[\hat{\beta}_n]$

Aside: $\widehat{\text{Avar}}[\hat{\beta}_n] = \underbrace{\left(\hat{\Sigma}_x + E_x \right)^{-1}}_{\hat{\Sigma}_x} \left(\Sigma_g + E_g \right) \left(\hat{\Sigma}_x + E_x \right)^{-1}$, where $E_x \equiv \hat{\Sigma}_x - \Sigma_x$
 $E_g \equiv \hat{\Sigma}_g - \Sigma_g$

CLT: $\sqrt{n}(\hat{\Sigma}_x - \Sigma_x) \xrightarrow{d} N(0, \text{some variance})$

intuitively: $\hat{\Sigma}_x - \Sigma_x = \frac{Z}{\sqrt{n}}$, where $Z \sim N(0, \text{some var})$

$$= \left(\Sigma_x + \frac{Z_x}{\sqrt{n}} \right)^{-1} \left(\Sigma_g + \frac{Z_g}{\sqrt{n}} \right) \left(\Sigma_x + \frac{Z_x}{\sqrt{n}} \right)^{-1}$$

(by Taylor's Thm)

$$\widehat{\text{Avar}}[\hat{\beta}_n] - \text{Avar}[\hat{\beta}_n] = \left(\Sigma_x + \frac{Z_x}{\sqrt{n}} \right)^{-1} \left(\Sigma_g + \frac{Z_g}{\sqrt{n}} \right) \left(\Sigma_x + \frac{Z_x}{\sqrt{n}} \right)^{-1} - \Sigma_x^{-1} \Sigma_g \Sigma_x^{-1}$$

CMT $\approx \left(\Sigma_x^{-1} + \frac{Z_x'}{\sqrt{n}} \right) \left(\Sigma_g + \frac{Z_g}{\sqrt{n}} \right) \left(\Sigma_x^{-1} + \frac{Z_x'}{\sqrt{n}} \right) - \Sigma_x^{-1} \Sigma_g \Sigma_x^{-1}$

$$O\left(\frac{1}{\sqrt{n}}\right) \rightarrow \approx \frac{Z_x'}{\sqrt{n}} \Sigma_g \Sigma_x^{-1} + \Sigma_x^{-1} \frac{Z_g}{\sqrt{n}} \Sigma_x^{-1} + \Sigma_x^{-1} \Sigma_g \frac{Z_x'}{\sqrt{n}}$$

$$O\left(\frac{1}{n}\right) \rightarrow + \frac{Z_x'}{\sqrt{n}} \cdot \frac{Z_g}{\sqrt{n}} \cdot \Sigma_x^{-1} + \frac{Z_x'}{\sqrt{n}} \Sigma_g \frac{Z_x'}{\sqrt{n}} + \Sigma_x^{-1} \frac{Z_x'}{\sqrt{n}} \frac{Z_x'}{\sqrt{n}}$$

$$O\left(\frac{1}{n^{3/2}}\right) \rightarrow + \frac{Z_x'}{\sqrt{n}} \frac{Z_g}{\sqrt{n}} \frac{Z_x'}{\sqrt{n}}$$

Condition (conditional homoskedasticity): $E[e_i^2 | x_i] = \sigma^2$ for all i

Under this condition, $\Sigma_g = E[e_i^2 x_i x_i^T]$

$$= E[E[e_i^2 | x_i] x_i x_i^T] \quad (\text{tower property})$$

$$\geq E[\sigma^2 x_i x_i^T]$$

$$= \sigma^2 \Sigma_x$$

$$\text{Avar}[\hat{\beta}_n] = \Sigma_x^{-1} \Sigma_g \Sigma_x^{-1} = \Sigma_x^{-1} (\sigma^2 \Sigma_x) \Sigma_x^{-1}$$

$\sigma^2 \Sigma_x^{-1}$ under conditional homoskedasticity

Recall we estimated Σ_x with $\hat{\Sigma}_x = \frac{1}{n} \sum_{i=1}^n x_i x_i'$

we estimate σ^2 with $s_n^2 = \frac{1}{n} \sum_{i=1}^n \hat{G}_i^2$

s_n^2 is consistent b/c $s_n^2 = \frac{1}{n} \sum_{i=1}^n \hat{G}_i^2 = \frac{1}{n} \sum_{i=1}^n (\epsilon_i + x_i' (\beta_0 - \hat{\beta}_n))^2$

$$= \underbrace{\frac{1}{n} \sum_{i=1}^n \hat{G}_i^2}_{\rightarrow \sigma^2} + \underbrace{\frac{2}{n} \sum_{i=1}^n \epsilon_i x_i' (\beta_0 - \hat{\beta}_n)}_{\rightarrow 0} + \underbrace{\frac{1}{n} \sum_{i=1}^n (\beta_0 - \hat{\beta}_n) x_i x_i' (\beta_0 - \hat{\beta}_n)'}_{\rightarrow 0}$$

Under conditional homoskedasticity: $\widehat{\text{Avar}}[\hat{\beta}_n] = s_n^2 \left(\frac{1}{n} \sum_{i=1}^n x_i x_i' \right)^{-1}$

$$z\text{-statistic} = \frac{\sqrt{n} [\hat{\beta}_n]_j}{\left[\widehat{\text{Avar}}[\hat{\beta}_n] \right]_{jj}^{1/2}} = \frac{\sqrt{n} [\hat{\beta}_n]_j}{\left[s_n^2 \left(\frac{1}{n} X'X \right)^{-1} \right]_{jj}^{1/2}}$$

$$= \frac{\hat{\beta}_n}{s_n \left[(X'X)^{-1} \right]_{jj}^{1/2}}$$