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"A spatial model incorporating dynamic, endogenous network interdependence: a political science application"

Web Appendix

Contents:

Expanded References

(The list below follows the order of the topics mentioned in the introduction of Franzese, Hays, & Kachi 2009. The complete reference list appears at the end of this Appendix.)

A Subject-Organized Reference List for Applied Spatial-Modeling in Political Science:

On Policy-Innovation Diffusion Among US States: Crain 1966; Walker 1969, 1973; Gray 1973; Knoke 1982; Caldiera 1985; Lutz 1987; Berry & Berry 1990; Case et al. 1993; Berry 1994; Rogers 1995; Mintrom 1997ab; Brueckner 1998; Mintrom & Vergari 1998; Mossberger 1999; Berry & Berry 1999; Godwin & Schroedel 2000; Balla 2001; Mooney 2001; Wejnert 2002; Coughlin et al. 2003; Bailey & Rom 2004; Boehmke & Witmer 2004; Daley & Garand 2004; Grossback et al. 2004; Mencken 2004; Berry & Baybeck 2005; Garrett et al. 2005; Costa-Font & Ons-Novell 2006; Karch 2006; Rincke 2006; Shipan & Volden 2006; Volden 2006; Werck et al. 2006; Woods 2006; Volden et al. 2007.

On Inter/Cross-National Policy-Innovation Diffusion: Schneider & Ingram 1988; Rose 1993; Bennett 1997; Dolowitz & Marsh 2000; True & Mintrom 2001; Tews et al. 2003; Jensen 2004; Meseguer 2004, 2005; Brooks 2005, 2007; Gilardi 2005; Gilardi et al. 2005; Murillo & Schrank 2005; Weyland 2005; Braun & Gilardi 2006; Linos 2006; Parys 2006; Ermini & Santolini 2007; Moscone et al. 2007.

On Institutional/Regime Diffusion: Dahl's 1971 classic *Polyarchy*, e.g., implicitly references international interdependence among the eight causes of democracy he lists; Starr's 1991 "Democratic Dominoes" and Huntington's 1991 *Third Wave* accord it a central role; Beissinger 2007 and Bunce & Wolchik 2006, 2007, *inter alia*, emphasize it in the context of post-communist democratic transitions in Eastern Europe, and Hagopian & Mainwaring 2005 among others in the Latin American context; finally, O'Loughlin et al. 1998, Brinks & Coppedge 2006, and Gleditsch & Ward 2006, 2007 estimated empirically the extent, paths, and/or patterns of international diffusion of democracy. Kelejian et al. 2007 give institutional diffusion general theoretical and empirical treatment.

Empirical Attention to the Inherent Interdependence of International

Relations… is most extensive in the work of Ward, Gleditsch, and colleagues—Shin & Ward 1999; Gleditsch & Ward 2000; Gleditsch 2002; Ward & Gleditsch 2002; Hoff & Ward 2004; Gartzke & Gleditsch 2006; Salehyan & Gleditsch 2006; Gleditsch 2007—and, in a different way, in Signorino and colleagues— Signorino 1999, 2002, 2003; Signorino & Yilmaz 2003; Signorino & Tarar 2006.

On Diffusion in Comparative & International Political Economy, and Globalization: Simmons & Elkins 2004 and Simmons et al. 2006, e.g., stress cross-national diffusion as the main force behind recent economic liberalizations, as do Eising 2002; Brune et al. 2004; Brooks 2005, 2007; Jordana & Levi-Faur 2005; Way 2005; Lazer 2006; Prakash & Potoski 2006; Brune & Guisinger 2007; and many others. Empirical work on globalization-induced interdependencies are far too numerous even to cite. Just a list of recent works emphasizing those that recognize explicitly that interdependence implies effects of some units outcomes on others—and still a small subset at that—would include Genschel 2002; Guler et al. 2002; Hays 2003; Franzese & Hays 2003, 2004ab, 2005ab, 2006abc, 2007abcd, 2008abcd, 2009abc; Badinger et al. 2004; Basinger & Hallerberg 2004; Hays et al. 2005; Heichel et al. 2005; Henisz et al. 2005; Holzinger & Knill 2005; Knill 2005; Polillo & Guillén 2005; Elkins et al. 2006; Jahn 2006; Lee & Strang 2006; Manger 2006; Swank 2006; Baturo & Grey 2007; Cao 2007; Cao et al. 2007; Coughlin et al. 2007; Garretsen & Peeters 2007; Mosley & Uno 2007; Mukherjee & Singer 2007.

On Interdependence of Legislators' Votes (Modeled Spatially): See, for example, Lacombe & Shaughnessy 2005.

On Interdependence of Citizens' Votes (Modeled Spatially): See, for example, Huckfeldt & Sprague 1991; O'Laughlin et al. 1994; Pattie & Johnston 2000; Beck et al. 2003; Calvo & Escolar 2003; Kim et al. 2003; Schofield et al. 2003; Lacombe & Shaughnessy 2007.

On Interdependence of Election Outcomes (Modeled Spatially): See, for example, Shin & Agnew 2002, 2007; Hiskey & Canache 2005; Wing & Walker 2006; Kayser 2007.

On Interdependence of Candidate Qualities, Contributions, or Strategies: See, for example, Goldenberg et al. 1986; Mizruchi 1989; Krasno et al. 1994; Cho 2003; Gimpel et al. 2006.

For Spatial Models of the Interdependence of the probabilities and outcomes of coups: e.g., Li & Thompson 1975; Of Riots: e.g., Govea & West 1981; Of Civil Wars: e.g., Murdoch & Sandler 2004, Buhaug & Rød 2006; Of Revolutions: e.g., Brinks & Coppedge 2006.

On Interdependence in Treaty Signing: see, e.g., Murdoch et al. 2003.

On Interdependence in Terrorist Origins and Targets: see, e.g., Brathwaite & Li 2008.

Contextual/Neighborhood Effects in Micro-Behavioral Studies: Huckfeldt & Sprague (1993) review the large literature on contextual/neighborhood effects in political behavior; as do Sampson et al. (2002) and Dietz (2002) for sociology. Recent analyses that stress interdependence include Straits 1990; O'Loughlin et al. 1994; Knack & Kropf 1998; Liu et al. 1998; Braybeck & Huckfeldt 2002ab; Beck et al. 2002; McClurg 2003; Huckfeldt et al. 2005; Cho & Gimpel 2007; Cho & Rudolph 2007.

On Interdependent Social-Movements: see, e.g., McAdam & Rucht 1993; Conell & Cohn 1995; Giugni 1998; Strang & Soule 1998; Biggs 2003; Browning et al. 2004; Andrews & Biggs 2006; Holmes 2006; Swaroop & Morenoff 2006.

On Interdependence in Violence and Crime: see, e.g., Grattet et al. 1998; Myers

2000; Baller et al. 2001; Morenoff et al. 2001; Villareal 2002; Baker & Faulkner 2003; Oberwittler 2004; Bhati 2005ab; Mears & Bhati 2006.

On Interdependence in (Microeconomic) Utilities: see, e.g., Akerloff 1997; Postlewaite 1998; Glaeser & Scheinkman 2000; Manski 2000; Brock & Durlauf 2001; Durlauf 2001; Glaeser et al. 2003; Yang & Allenby 2003; Sobel 2005; Ioannides 2006; Soetevent 2006.

On Interdependence in Macroeconomic Performance: see, e.g., Fingleton 2003; Novo 2003; Kosfeld & Lauridsen 2004; Maza & Villaverde 2004; Kelejian et al. 2006; Mencken et al. 2006.

On Interdependence in Technology, Marketing, and other Firm Strategies: see, e.g., ; Abramson & Rosenkopf 1993; Geroski 2000; Strang & Macy 2001; Holloway 2002; Bradlow 2005; Autant-Berard 2006; Mizruchi et al. 2006.

On Interdependence in Fertility, Birthweight, Child development, or Child Poverty: see, e.g., Tolnay 1995, Montgomery & Casterline 1996; Morenoff 2003; Sampson et al. 1999; Voss et al. 2006.

On Interdependence in ordainment of women: Chaves 1996; in right-wing extremism: Rydgren 2005, in marriage: Yabiku 2006, in (sub)national identity: Lin et al. 2006; in obesity: Christakis & Fowler 2007; and in research faculty: Weinstein 2007.

The intellectual-historical genesis of *Galton's Problem*:

Galton originally raised the issue thus: "[F]ull information should be given as to the degree in which the customs of the tribes and races which are compared together are independent. It might be that some of the tribes had derived them from a common source, so that they were duplicate copies of the same original. ...It would give a useful idea of the distribution of the several customs and of their relative prevalence in the world, if a map were so marked by shadings and colour as to present a picture of their geographical ranges" (Galton 1889, as quoted in Darmofal 2007). We find further historical context in http://en.wikipedia.org/wiki/Galton's_problem: "In [1888], Galton was present when Sir Edward Tylor presented a paper at the Royal Anthropological Institute. Tylor had compiled information on institutions of marriage and descent for 350 cultures and examined the correlations between these institutions and measures of societal complexity. Tylor interpreted his results as indications of a general evolutionary sequence, in which institutions change focus from maternal to paternal lines as societies grow more complex. Galton disagreed, noting that similarity between cultures could be due to borrowing, could be due to common descent, or could be due to evolutionary development; he maintained that without controlling for borrowing and common descent one cannot make valid inferences regarding evolutionary development. Galton's critique has become the eponymous *Galton's Problem* (Stocking 1968:175), as named by Raoul Naroll (1961, 1965), who proposed [some of] the first statistical solutions."

Spatial-econometric and spatial-statistical approaches to spatial analysis:

Methodologically, two approaches to spatial analysis can be discerned: spatial statistics and spatial econometrics (Anselin 2002, 2006). The distinction rests, on the one hand, on the relative emphasis in spatial-econometric approaches to theoretical models of interdependence processes wherein space may often have broad meaning, well beyond geography and geometry to encompass all manner of social, economic, or political connection that induces effects from outcomes in some units on outcomes in others (Brueckner 2003; Beck et al. 2006). The spatial-lag regression model plays a starring role in that tradition (Hordijk 1974; Paelinck and Klaassen 1979; Anselin 1980, 1988, 1992; Haining 1990; LeSage 1999). According to Anselin, theory driven models deal with substantive spatial correlation (2002). This approach has its own method to model specification and estimation. The importance of spatial interdependence is tested using Wald tests and the unrestricted spatial lag model. On the other hand, spatial-error models, analysis of spatial-correlation patterns, spatial kriging, and spatial smoothing, e.g., characterize the more-exclusively data-driven approaches and the typically narrower conception of space in solely geographic/geometric terms in the longer spatial-statistics tradition (initially inspired by Sir Galton's famous comments at the 1888 meetings of the Royal Anthropological Society, and reaching crucial methodological milestones in Whittle 1954; Cliff & Ord 1973, 1981; Besag 1974; Ord 1975; Ripley 1981; Cressie 1993). According to Anselin, this kind of modeling is driven by data problems such as measurement error. The spatial correlation is viewed as a *nuisance*. This approach has its own method to model specification and estimation. The presence of spatial effects is tested using Lagrange multiplier tests and the restricted non-spatial lag model.

Anselin (2002, 2006), Griffith & Paelinck (2007), and many others discern two approaches to spatial analysis: spatial statistics and spatial econometrics. "In practice, spatial econometrics and spatial statistics reflect traditions of their parent disciplines. In other words, they share much in common, with some notably differing emphases… Econometrics…is the 'setting up of mathematical models describing economic relationships…, testing the validity of such hypotheses and estimating the parameters in order to obtain a measure of the strengths of the influences of the different independent variables' (Bannock et al. 2003)… 'Statistics is the science of gaining information from numerical data' (Moore 1995, p. 2). It provides data-interrogative tools and conceptual

frameworks for gaining understanding through empirical-based induction, and involves data acquisition, data analysis, and statistical inference" (Griffith & Paelinck 2007:210). Thus, the distinction rests, on the one hand, on the relative emphasis in spatial-econometric deductive approaches to theoretical models of interdependence processes, wherein space may often have broad meaning, well beyond physical distance to encompass all manner of social, economic, or political connection that induces effects from outcomes in some units on outcomes in others (see, e.g., recently, from economics and political science: Brueckner 2003; Beck et al. 2006). Spatial-lag models play starring roles in that tradition (Hordijk 1974; Paelinck & Klaassen 1979; Anselin 1980, 1988, 1992; LeSage 1999). According to Anselin (2002), such theory-driven models deal with substantive spatial correlation, and this approach lends itself to methods of model specification, estimation, and evaluation that begin with the unrestricted spatial lag model and use Wald-style testing to gauge the importance of spatial interdependence. On the other hand, spatial-error models, analysis of spatial-correlation patterns, spatial kriging, and spatial filtering, for examples, characterize the more-exclusively data-driven, i.e., inductive, approaches and the more-typically narrower conception of space in solely geographic/geometric terms in the longer spatial-statistics tradition (initially inspired by Sir Galton's famous comments at the 1888 meetings of the Royal Anthropological Society, and reaching methodological milestones in Whittle 1954; Cliff & Ord 1973, 1981; Besag 1974; Ord 1975; Ripley 1981; Cressie 1993). Again following Anselin (2002), this modeling approach stresses spatial patterns and clustering, seen more often as driven by data problems such as measurement error, and so tends to treat spatial correlation more as nuisance than substance. Such emphases lend themselves more naturally to the opposite direction in model specification, estimation, and evaluation, beginning with the restricted non-spatial model and adopting Lagrange-multiplier-style testing.[1] We would mention additionally only how the core questions asked and sorts of answers sought tend to differ across approaches. The core substantive distinction drawn in our proposal between spatial association arising from common exposure, contagion, or selection, and the challenges of separately identifying and estimating them are typically more centrally stressed from spatial-econometric perspectives (model specified

[1] A debate chronicled in *Regional Science and Urban Economics* (Florax et al. 2003, 2006; Hendry 2006; Mur & Angulo 2009) continues as to whether one can demonstrate inferential dominance of the bottom-up or top-down strategy.

deductively for purposes of inference about processes) than from spatial-statistical ones. Indeed, the fact of association will often suffice from a spatial-statistical vantage (inductive process of gaining information from the data), so the modeling approach often makes no attempt to distinguish them. Likewise, the kinds of counterfactuals of interest to us arise more naturally in a spatial-econometric than a spatial-statistical framework.

Network analysis, finally, considers a related (but not identical) set of substantive questions regarding the structure of ties (edges, arcs, connections) between units (nodes, actors, vertices). Central questions concern characterizing (measuring) the structure of networks, explaining their genesis, and, less centrally, at least until very recently, considering the effects of network structure and of units' location within it on units' actions (behavior, attributes). Again until perhaps recently, questions of how other units' (alters') actions affect each unit (ego) via the connections given by the network seem to have been less central still. Methods for network analysis seem to have originated primarily in sciences (physics, biology, computer-science) and mathematics, rather than in statistics or econometrics, and methodological development, including the eventual importation of statistical concepts, theories, and methods, has been largely separate from either spatial statistics or spatial econometrics (again, until recently perhaps), notwithstanding the similarity of their substantive interests.[2] The typical core questions of network analysis, thus, are subtly distinct from those of spatial econometrics and spatial statistics. Our own emphasis on Galton's Problem, the distinguishing of common exposure from interdependence (a.k.a., contagion) sources of spatial association, takes an important additional concern from a core network-analysis question of explaining network genesis, i.e., selection (a.k.a., homophily).

[2] Consider, e.g., that the canonical reference in social-network analysis, Wasserman & Faust (1994) does not cite Cliff & Ord, Cressie, or Anselin, and that Carrington et al. (2005), intended as an update and companion to Wasserman & Faust cites only Cliff & Ord (1981) on just two pages of just one section and only with regard to Moran's I. Getis et al.'s (2004) *Spatial Econometrics and Spatial Statistics*, symmetrically, cites neither Wasserman & Faust nor Carrington et al.

Monte Carlo Simulation of S-ML vs. S-OLS vs. OLS Estimation of m-STAR Models

Before illustrating the estimation, testing, and interpretation of our m-STAR model of network-behavior coevolution, we will demonstrate that, in fact, the S-ML estimators just described are needed and outperform simpler least-squares estimators. Analytically, the omitted-variable biases of the *blind* OLS estimator remain as before: $\mathbf{F\beta}$. The simultaneity asymptotic biases (inconsistencies) of the *naïve* S-OLS estimator, which simply inserts the multiple spatial lags into least-squares regression, are also analogous to **Error! Reference source not found.** as follows:

$$\text{Let } \mathbf{y} = \mathbf{Z\rho} + \boldsymbol{\varepsilon} \;\; ; \;\; \text{where } \mathbf{Z} = \begin{bmatrix} \mathbf{W}_1\mathbf{y} & \mathbf{W}_2\mathbf{y} \end{bmatrix} \text{ and } \boldsymbol{\rho} = \begin{bmatrix} \rho_1 & \rho_2 \end{bmatrix}'.$$

$$\text{Then: plim } \hat{\boldsymbol{\rho}} = \boldsymbol{\rho} + \text{plim}\left[\left(\frac{\mathbf{Z}'\mathbf{Z}}{n}\right)^{-1}\frac{\mathbf{Z}'\boldsymbol{\varepsilon}}{n}\right]; \text{ that is:}$$

$$\text{plim } \hat{\boldsymbol{\rho}} = \begin{bmatrix} \rho_1 \\ \rho_2 \end{bmatrix} + \begin{bmatrix} \dfrac{V\left(\mathbf{W}_2\mathbf{y}\right)}{V\left(\mathbf{W}_2\mathbf{y}\right)\times V\left(\mathbf{W}_1\mathbf{y}\right) - \left[C\left(\mathbf{W}_1\mathbf{y},\mathbf{W}_2\mathbf{y}\right)\right]^2} & \dfrac{-C\left(\mathbf{W}_1\mathbf{y},\mathbf{W}_2\mathbf{y}\right)}{V\left(\mathbf{W}_2\mathbf{y}\right)\times V\left(\mathbf{W}_1\mathbf{y}\right) - \left[C\left(\mathbf{W}_1\mathbf{y},\mathbf{W}_2\mathbf{y}\right)\right]^2} \\ \dfrac{-C\left(\mathbf{W}_1\mathbf{y},\mathbf{W}_2\mathbf{y}\right)}{V\left(\mathbf{W}_2\mathbf{y}\right)\times V\left(\mathbf{W}_1\mathbf{y}\right) - \left[C\left(\mathbf{W}_1\mathbf{y},\mathbf{W}_2\mathbf{y}\right)\right]^2} & \dfrac{V\left(\mathbf{W}_1\mathbf{y}\right)}{V\left(\mathbf{W}_2\mathbf{y}\right)\times V\left(\mathbf{W}_1\mathbf{y}\right) - \left[C\left(\mathbf{W}_1\mathbf{y},\mathbf{W}_2\mathbf{y}\right)\right]^2} \end{bmatrix} \begin{bmatrix} C\left(\boldsymbol{\varepsilon},\mathbf{W}_1\mathbf{y}\right) \\ C\left(\boldsymbol{\varepsilon},\mathbf{W}_2\mathbf{y}\right) \end{bmatrix}; \quad (1).$$

$$\text{that is: plim } \hat{\boldsymbol{\rho}} = \begin{bmatrix} \rho_1 \\ \rho_2 \end{bmatrix} + \begin{bmatrix} \dfrac{\Psi_{22}\Gamma_{11} - \Psi_{12}\Gamma_{21}}{|\boldsymbol{\Psi}|} \\ \dfrac{\Psi_{11}\Gamma_{21} - \Psi_{12}\Gamma_{11}}{|\boldsymbol{\Psi}|} \end{bmatrix}, \text{ where } \boldsymbol{\Psi} = \text{plim}\left(\frac{\mathbf{Z}'\mathbf{Z}}{n}\right) \text{ and } \boldsymbol{\Gamma} = \text{plim}\left(\frac{\mathbf{Z}'\boldsymbol{\varepsilon}}{n}\right)$$

The intuitions remain as before: simultaneity biases generally increase in ρ, concentrating in the spatial lags that covary most with the residual, and inducing biases in generally opposite directions for other spatial-lag and covariates' coefficient-estimates.

To demonstrate that an estimator is inconsistent, however, does not demonstrate that these asymptotic biases are practically large or that they outweigh other potential deficiencies of consistent estimators. Accordingly, we conduct some simple Monte Carlo simulations to explore the small-sample performance of these estimators, specifically the magnitudes of their biases, inefficiency, and standard-error inaccuracy. With the analytical results in (1), Table 1's simulation results are easily interpreted. The covariance of the *queen* spatial-lag (all eight adjacent squares on a grid) and $\boldsymbol{\varepsilon}$ is about half that of the *rook* lag (only the four horizontally and vertically adjacent) and $\boldsymbol{\varepsilon}$. With row-standardization, eight ties, and fixed ρ, the strength of the interdependence/endogeneity is more diluted in *queen*-lag $\mathbf{W}_2\mathbf{y}$. Consequently, $\hat{\rho}_1$ on *rook*-lag $\mathbf{W}_1\mathbf{y}$ is overestimated, and $\hat{\rho}_2$ on *queen*-lag $\mathbf{W}_2\mathbf{y}$ and $\hat{\beta}_0$ on x_0 (which is

especially correlated with $\mathbf{W}_1\mathbf{y}$)[3] are (badly) underestimated. The S-ML estimator also dominates impressively in efficiency and standard-error accuracy, especially for those three coefficients.

Table 1. Estimator Comparison for m-STAR Model: S-OLS vs. S-ML

ESTIMATOR	RESULT	$\boldsymbol{\beta_0}$**=1**	$\boldsymbol{\beta_1}$**=1**	$\boldsymbol{\rho_1}$**=.3**	$\boldsymbol{\rho_2}$**=.3**
S-OLS	*Average Estimate*	.38 / .24	.96 / .97	.47 / .47	.27 / .29
	Standard Deviation	.33 / .24	.07 / .05	.19 / .15	.21 / .16
	Root Mean-Squared-Error	.71 / .80	.08 / .06	.25 / .22	.21 / .16
	Average Std-Err Estimate	.37 / .28	.06 / .05	.15 / .11	.17 / .13
	Overconfidence	.92 / .87	1.06 / 1.02	1.29 / 1.35	1.21 / 1.27
S-ML	*Average Estimate*	1.09 / 1.05	1.00 / 1.00	.31 / .31	.27 / .28
	Standard Deviation	.33 / .24	.07 / .05	.12 / .09	.14 / .11
	Root Mean-Squared-Error	.34 / .24	.07 / .05	.12 / .09	.14 / .11
	Average Std-Err Estimate	.31 / .23	.06 / .05	.12 / .09	.14 / .11
	Overconfidence	1.05 / 1.05	1.03 / 1.00	.98 / 1.01	.98 / 1.01

Monte Carlo (1000 Trials) Results for $\mathbf{y}=\rho_1\mathbf{W}_1\mathbf{y}+\rho_2\mathbf{W}_2\mathbf{y}+\mathbf{X}\boldsymbol{\beta}+\boldsymbol{\varepsilon}$, with $\mathbf{W}_1$=rook adjacency, $\mathbf{W}_2$=queen adjacency (row normalized); $\beta_0=\beta_1=1$, $\rho_0=\rho_1=.3$; and N=225/450.

Table 2 similarly evaluates the *blind* OLS, *naïve* S-OLS, and S-ML estimators for our m-STAR coevolution model. S-ML again outperforms the inconsistent OLS alternatives. Notice that, with $\mathbf{x}$ drawn independently, appreciable correlation of the regressors with the spatial lags concentrates in the time-lag; thus, omitted-variable biases of *blind* OLS are not severe for β and concentrate at a noticeable 20% in ϕ. Notice also that with endogenous $\mathbf{Ly}$ being time-lagged as estimated and in truth, and $\mathbf{L}$ being $|y_i\text{-}y_j|$ and so not terribly linearly related to y_i, the simultaneity biases of S-OLS concentrate in ρ at a sizable +33%, whereas $\mathbf{x}$ being drawn independently, little induced bias appears in $\hat{\beta}$. The efficiency (RMSE) and standard-error accuracy gains are more uniform and obvious.

[3] The simulations drew $\boldsymbol{\varepsilon}$ from N(1,1), making the nonzero aspect of $\mathbf{Z}'\mathbf{Wy}$ concentrate in the constant, $\mathbf{x}_0$.

Table 2. Estimator Comparison for m-STAR Coevolution Model: OLS v. S-OLS v. S-ML

ESTIMATOR	RESULT	ϕ=.3	β=1	ρ=.3	γ=.3
OLS	*Average Estimate*	.36	1.04	—	—
	Standard Deviation	.07	.05	—	—
	Root Mean-Squared-Error	.09	.06	—	—
	Average Std-Err Estimate	.03	.05	—	—
	Overconfidence	2.07	.97	—	—
S-OLS	*Average Estimate*	.28	.99	.41	.25
	Standard Deviation	.03	.05	.06	.08
	Root Mean-Squared-Error	.04	.05	.129	.10
	Average Std-Err Estimate	.03	.04	.05	.08
	Overconfidence	1.02	1.05	1.19	1.01
S-ML	*Average Estimate*	.29	1.00	.31	.27
	Standard Deviation	.03	.05	.05	.08
	Root Mean-Squared-Error	.03	.05	.05	.09
	Average Std-Err Estimate	.03	.04	.04	.07
	Overconfidence	1.02	1.04	1.11	1.10

Monte Carlo (1000 Trials) Results for $\mathbf{y}_t=\rho\mathbf{W}\mathbf{y}_t+\gamma\mathbf{L}\mathbf{y}_{t-1}+\phi\mathbf{y}_{t-1}+\beta\mathbf{x}_t+\boldsymbol{\varepsilon}_t$, with $\mathbf{W}$=48 contiguous US-state adjacency pattern (row-stdzd); ρ=.3, γ=.3, ϕ=.3, β=1; and N=48, T=10.

In sum, even in simulations rather favorably designed for the *blind* or *naïve* estimators, the S-ML estimator is clearly dominant for all estimates and estimate-properties.

Additional Network-Estimates Graphs

Figure 4: The Estimated Network of ALM-policy Interdependence, 1981

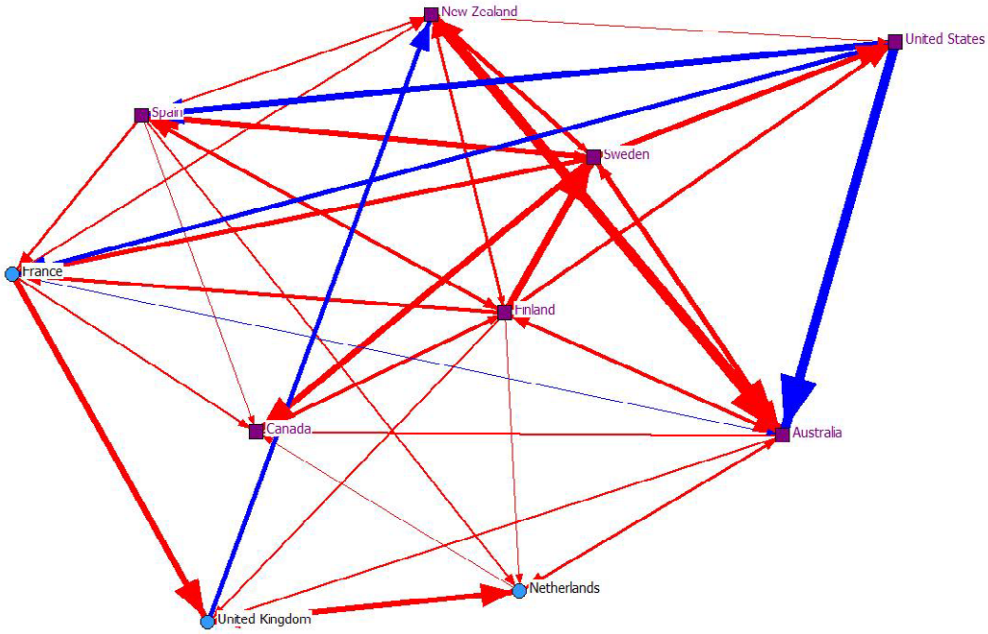
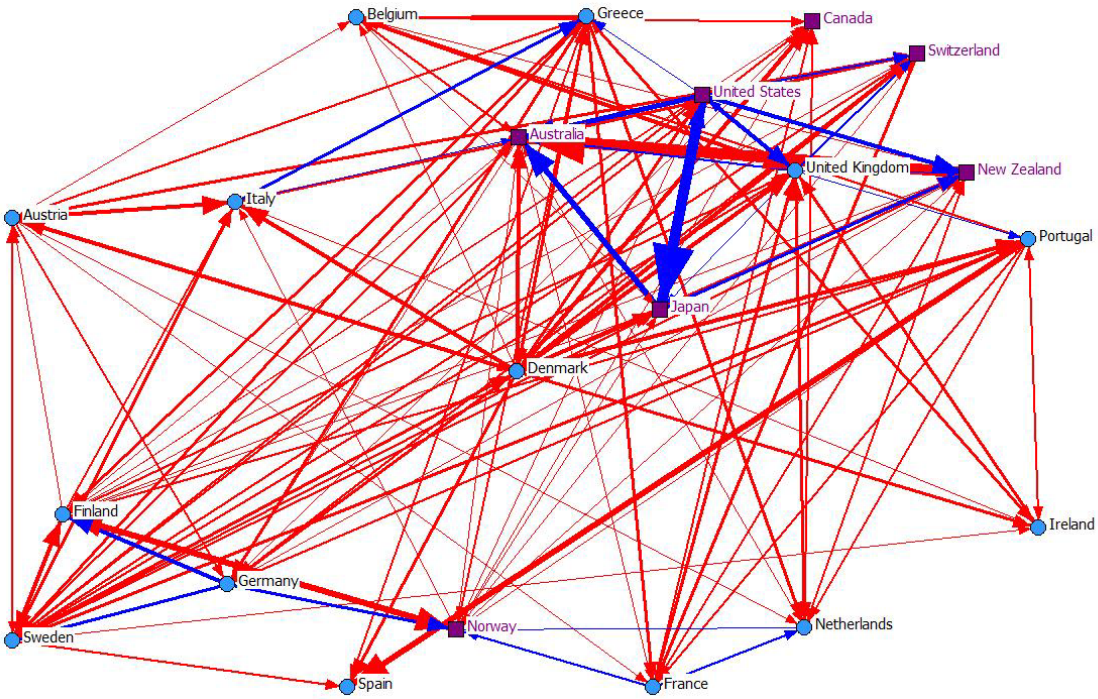


Figure 6: The Estimated Network of ALM-policy Interdependence, 2001



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