# A structural result for Personalized PageRank and its algorithmic consequences 

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## Motivation

Graphs arise in many domains


Questions that may help us understand graphs:
■ Which nodes are most important/influential, globally and locally?
■ Which nodes are similar/relevant to a given node?

One model to answer these questions: Personalized PageRank (PPR)

## PPR definition

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Given directed graph $G=(V, E)$, let $v \in V$ and $\alpha \in(0,1)$
Define Markov chain $\left\{X_{t}^{\vee}\right\}_{t \in \mathbb{N}}$ as follows: given $X_{t}^{\vee}$,
■ W.p. $(1-\alpha)$, sample $X_{t+1}^{v}$ from out-neighbors of $X_{t}^{v}$ (random walk)
■ W.p. $\alpha$, set $X_{t+1}^{v}=v($ jump to $v)$
Stationary distribution $\pi_{v}=\left\{\pi_{v}(w)\right\}_{w \in V}$ called PPR vector
Matrix $\Pi=\left\{\pi_{v}\right\}_{v \in V}$ called PPR matrix


## PPR interpretation

$\pi_{v}(w)$ large when $w$ frequently visited on short walks from $v$
$\Rightarrow$ Interpret $\pi_{\nu}(w)$ as measure of $w$ 's importance/relevance to $v$


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Example: suggest $v$ follow $w$ on Twitter if $\pi_{v}(w)$ large (Gupta et al. 2013)

- interested in technology, so
- PPR encodes this: $\pi_{v}\left(w_{1}\right) \gg \pi_{v}\left(w_{2}\right)$ when $v=w_{1}=w_{2}=0$



## PPR dimensionality

Using Perron-Frobenius theorem, can show $\operatorname{rank}(\Pi)=|V|=: n$

However, PPR exhibits transitive structure

- $\pi_{v_{1}}\left(v_{2}\right), \pi_{v_{2}}\left(v_{3}\right)$ large $\Rightarrow \pi_{v_{1}}\left(v_{3}\right)$ large ("friend of my friend is my friend")

■ Suggests $\Pi$ has small "effective dimension"

Also, for many real-world graphs $G=(V, E),|E|=O(n)$
■ Suggests $G$ is $O(n)$-dimensional, but $\Pi$ (derived from $G$ ) is $n^{2}$-dimensional
■ Why this gap? Is it actually present?

## Outline of talk:

1 How to quantify effective dimension of $\Pi$ ?
12 Can we bound this measure of dimensionality?
3 If bound "small", can we leverage it algorithmically?

## Quantifying PPR dimensionality

Natural measure of effective dimension of $\Pi$ :

$$
\begin{equation*}
\Delta(\epsilon)=\min _{\hat{\Pi}} \operatorname{rank}(\hat{\Pi}) \text { s.t. }\|\Pi-\hat{\Pi}\|<\epsilon \tag{1}
\end{equation*}
$$

Intuitively, $\Pi$ low dimensional if close to low-rank matrix

Can also view (1) as dual of low-rank approximation:

$$
\inf _{\hat{\Pi}}\|\Pi-\hat{\Pi}\| \text { s.t. } \operatorname{rank}(\hat{\Pi}) \leq k
$$

We take $\|\cdot\|=\|\cdot\|_{\infty}$ in (1), where for matrix $A$ with rows $a_{1}, \ldots, a_{n}$,

$$
\|A\|_{\infty}=\max _{i \in\{1, \ldots, n\}}\left\|a_{i}\right\|_{1}
$$

(Natural choice, since $\|\cdot\|_{T V}=\|\cdot\|_{1} / 2$ and each row of $\Pi$ is a distribution)

## Modified dimensionality measure

For analytical/algorithmic reasons, we let $K \subset V$ and upper bound $\Delta(\epsilon)$ as

$$
\begin{equation*}
\Delta(K, \epsilon)=|K|+\left|\left\{v \notin K: \min _{\mu_{v}(k)}\left\|\pi_{v}-\sum_{k \in K} \mu_{v}(k) \pi_{k}\right\|_{1}>\epsilon\right\}\right| \tag{2}
\end{equation*}
$$

- Think of $K$ as hub nodes (located "centrally" in graph)

■ Will argue that for most non-hubs, PPR close to linear combo of hub PPR

- Second term in (2) accounts for other non-hubs (typically "far" from hubs)



## Graph model

$\Delta(K, \epsilon)$ highly dependent on local graph structure - hard to bound in general
We analyze directed configuration model (DCM) due to "nice" local structure ${ }^{1}$

DCM construction:
1 Realize degree sequence $\left\{d_{\text {out }}(v), d_{\text {in }}(v)\right\}_{v \in V}$
$\sqrt{2}$ Attach $d_{\text {out }}(v)\left(d_{\text {in }}(v)\right.$, resp.) outgoing (incoming, resp.) half-edges to $v$
3 Randomly pair half-edges to form edges via breadth-first-search


[^0]
## Jump probability and dimensionality

Choice of $\alpha=\mathbb{P}($ jump to $v)$ impacts dimensionality:
■ $\alpha \approx 0 \Rightarrow \pi_{v} \approx$ random walk stationary distribution $\Rightarrow \Delta(K, \epsilon) \approx 1$

- $\alpha \approx 1 \Rightarrow \pi_{v} \approx$ point mass on $v \Rightarrow \Delta(K, \epsilon) \approx n$

How to make this precise?
Namely, for a sequence $\left\{G_{n}\right\}_{n \in \mathbb{N}}$ of DCMs, how should $\alpha=\alpha_{n}$ scale with $n$ ?


## Jump probability and mixing times

Suppose $\alpha_{n} \log n \rightarrow 0$, e.g. $\alpha_{n}=1 /(\log n)^{2}$
Since $\alpha_{n}=\mathbb{P}($ jump to $v), \mathbb{E}[$ random walk length $]=\Theta\left((\log n)^{2}\right)$
Bordenave, Caputo, Salez 2018: random walk on DCM mixes in $\Theta(\log n)$ steps
Mixing occurs before jump to $v$ ! Allows us to show $\Delta(K, \epsilon)=1$ with high prob.
Hence, we set $\alpha_{n}=\Theta(1 / \log n)$ (just outside the trivial regime)


## Main result

Main result concerns sequence of DCMs $\left\{G_{n}\right\}_{n \in \mathbb{N}}$, where $G_{n}$ has $n$ nodes

From $G_{n}$, we define $\Delta_{n}\left(K_{n}, \epsilon\right)$ (a random variable, since $G_{n}$ is random)

Our main result says $\Delta_{n}\left(K_{n}, \epsilon\right)=o(n)$ with high probability as $n \rightarrow \infty$ :

## Theorem

Assume degree sequence satisfies certain assumptions (details to come), and assume $\alpha_{n}=\Theta(1 / \log n)$. Then for any $\epsilon>0$, some $c_{\epsilon} \in(0,1)$, and any $C>0$, all independent of $n$,

$$
\lim _{n \rightarrow \infty} \mathbb{P}\left(\Delta_{n}\left(K_{n}, \epsilon\right)>C n^{c_{\epsilon}}\right)=0
$$

## Proof of main result

Main result follows almost immediately from key lemma:

## Lemma

Under assumptions of theorem, we have for $s \sim V$ uniformly and some $\tilde{c}_{\epsilon}>0$,

$$
\mathbb{P}(\min _{\mu_{s}(k)} \underbrace{\left\|\pi_{s}-\sum_{k \in K} \mu_{s}(k) \pi_{k}\right\|_{1}}_{*}>\epsilon)=O\left(n^{-\tilde{\tau}_{\epsilon}}\right) .
$$

Outline for proof of lemma:
1 Show $\star$ depends only on neighborhood of $s$ for certain $\mu_{s}(k)$
$\sqrt{2}$ Approximate neighborhood construction with branching process (using Chen, Litvak, Olvera-Cravioto 2017) to study $\star$ on tree
3 Recursive nature of branching process $\rightarrow \star$ on tree is martingale-like $\rightarrow$ analyze similar to method of bounded differences

## Choice of $\mu_{v}(k)$

By considering first step of PPR Markov chain, can show

$$
\pi_{v}(w)=\underbrace{\alpha 1(w=v)}_{\text {first step is jump to } v}+\underbrace{\sum_{k: v \rightarrow k} \frac{(1-\alpha)}{|\{k: v \rightarrow k\}|} \pi_{k}(w)}_{\text {first step follows random walk }}
$$

For any $K \subset V$, Jeh, Widom 2003 proves decomposition of same form:

$$
\pi_{v}(w)=\frac{\alpha 1(w \notin K) \tilde{\pi}_{v}(w)}{\alpha+(1-\alpha) \tilde{\pi}_{v}(K)}+\sum_{k \in K} \frac{\tilde{\pi}_{v}(k)}{\alpha+(1-\alpha) \tilde{\pi}_{v}(K)} \pi_{k}(w)
$$

where $\tilde{\pi}_{v}$ is PPR on graph with outgoing edges from $K$ removed

In proof (and later, in algorithm), we let $\mu_{v}(k)=\frac{\tilde{\pi}_{v}(k)}{\alpha+(1-\alpha) \tilde{\pi}_{v}(K)}$

## Assumptions (1/2)

Recall: DCM randomly pairs edges from degree sequence $\left\{d_{\text {out }}(v), d_{\text {in }}(v)\right\}_{v \in v}$


We assume $\left\{d_{\text {out }}(v), d_{\text {in }}(v)\right\}_{v \in V}$ satisfies two properties with high probability
Property 1: $\left\{d_{\text {out }}(v), d_{\text {in }}(v)\right\}_{v \in V}$ is sparse (e.g. $O(n)$ total edges)
$\Rightarrow$ Needed for branching process approximation; possible artifact of analysis
Property 2: $|K|=o(n)$ but $K$ contains non-vanishing fraction of edges, i.e.

$$
\frac{\sum_{k \in K} d_{i n}(k)}{\sum_{v \in V} d_{i n}(v)} \xrightarrow[n \rightarrow \infty]{ } p>0
$$

$\Rightarrow$ We believe this assumption is fundamentally necessary

## Assumptions (2/2)

Recall key property:

$$
\begin{equation*}
|K|=o(n), \quad \frac{\sum_{k \in K} d_{i n}(k)}{\sum_{v \in V} d_{i n}(v)} \xrightarrow[n \rightarrow \infty]{ } p>0 \tag{3}
\end{equation*}
$$

Empirically holds if $d_{\text {in }}(v)$ follow power law, common model for e.g. Twitter


## Geometric interpretation of theorem

Theorem says for most $v \notin K$ and some $\mu_{v}(k) \geq 0$,

$$
\pi_{v} \approx \sum_{k \in K} \mu_{v}(k) \pi_{k}
$$

When $|V|$ large, we also show $\sum_{k \in K} \mu_{v}(k) \approx 1$, so for most $v \notin K$,

$$
\pi_{v} \approx \text { convex combination of }\left\{\pi_{k}\right\}_{k \in K}
$$

$\Rightarrow$ Most of $\left\{\pi_{v}\right\}_{v \notin K}$ lie near convex hull of $\left\{\pi_{k}\right\}_{k \in K}$, which shrinks relative to $|V|$-dimensional simplex (a few $\left\{\pi_{v}\right\}_{v \notin K}$ can be far away)




## Empirical results (1/2)

Compute bound on $\left\|\pi_{v}-\sum_{k \in K} \mu_{v}(k) \pi_{k}\right\|_{1}$, averaged across $v \notin K$

Set $K=$ nodes of highest in-degree, $\alpha_{n}=1 / \log n$


For DCM with power law in-degrees, average error decays as $n$ grows (despite $|K| / n$ decaying too)


For variety of real graphs, average error decays as $\kappa$ grows when $K=$ $n^{\kappa}$ nodes of highest in-degree

## Empirical results (2/2)

Bound $\Delta(K, \epsilon)$ for two real graphs (social network, partial web crawl)
$K$ and $\alpha_{n}$ chosen as in previous slide
For soc-Pokec, $\Delta(K, \epsilon)=0.09 n$ when $\epsilon=\frac{1-\alpha_{n}}{3}$; similar for web-Google ${ }^{2}$
Thus, while theorem doesn't apply, $\Delta(K, \epsilon)$ small relative to $n$ for reasonable $\epsilon$


[^1]
## Baseline algorithm (Jeh, Widom 2003)

Jeh, Widom 2003 proposes (but doesn't analyze!) the following:
1 Choose "hub" nodes, estimate PPR vectors directly
$\boxed{2}$ For other nodes, estimate PPR as linear combo of hub $\mathrm{PPR}^{3}$
Our result $\Rightarrow$ linear combo good estimate for all but $O(n)$ non-hubs if $O(n)$ hubs
Thus, we improve Jeh, Widom 2003, but questions remain:
■ Can we guarantee accuracy all nodes?
■ Can we estimate hub PPR, and non-hub linear combo weights, with provably good performance? (only heuristics in Jeh, Widom 2003)


[^2]
## Improving accuracy of baseline scheme

Baseline scheme: for $v \notin K, \pi_{v}$ estimated as

$$
\hat{\pi}_{v}=\sum_{k \in K} \mu_{v}(k) \pi_{k}
$$

where $\mu_{v}(k)$ from linear decomposition shown previously

We show (for a certain function $f$ )

$$
\left\|\pi_{v}-\hat{\pi}_{v}\right\|_{1}<\epsilon \Leftrightarrow \sum_{k \in K} \mu_{v}(k)>f(\epsilon)
$$

Intuitively, small error $\Leftrightarrow v$ is "close" to $K$ in graph

Key point: $\sum_{k \in K} \mu_{v}(k)$ is (approximately) known at runtime!

$$
\Rightarrow \text { If } \sum_{k \in K} \mu_{v}(k)<f(\epsilon) \text {, estimate } \pi_{v} \text { directly }
$$

## Estimating PPR and linear combo weights (1/2)

Recall: $\pi_{v}=$ stationary distribution of chain with transition matrix

$$
P_{v}=\underbrace{(1-\alpha) P}_{\text {Random walk }}+\underbrace{\alpha 1_{n} e_{v}^{\top}}_{\text {Jump to } v}
$$

Solving $\pi_{v}=\pi_{v} P_{v}$ yields

$$
\pi_{v}=\alpha e_{v}^{\top}\left(I_{n}-(1-\alpha) P\right)^{-1}
$$

Since $\pi_{v}$ is $v$-th row of $\Pi$,

$$
\Pi=\alpha\left(I_{n}-(1-\alpha) P\right)^{-1}=\alpha \sum_{i=0}^{\infty}(1-\alpha)^{i} P^{i}
$$

Suggests power iteration: choose $i^{*}$ large and compute

$$
\alpha \sum_{i=0}^{i^{*}}(1-\alpha)^{i} P^{i} \approx \Pi
$$

## Estimating PPR and linear combo weights (2/2)

Power iteration traverses all paths of length $\leq i^{*}$

Dynamic programming (DP) variants traverse only "important" paths

Forward DP (Andersen, Chung, Lang 2006):
■ Given $v$, traverses "important" paths out of $v$; estimates $v$-th row of $\Pi$
■ Can use to estimate PPR vectors directly

Backward DP (Andersen et al. 2008):

- Given $v$, traverses "important" paths into $v$; estimates $v$-th column of $\Pi$

■ Can use (modified version) to estimate linear combo weights

## Putting it all together

Our scheme estimates $\pi_{v} \ldots$

- ... by forward DP, if $v \in K$

■ ... by forward DP, if $v \notin K$ and linear combo determined to be inaccurate
$■ \ldots$ as linear combo, if $v \notin K$ and linear combo determined to be accurate
Forward DP provably accurate; thus, all estimates are accurate

Complexity dominated by number runs of forward DP

- By design, forward DP is run $\Delta(K, \epsilon)$ times
- Each run has $O(n \log n)$ complexity (by Andersen, Chung, Lang 2006) ${ }^{4}$

Overall complexity is $O(\Delta(K, \epsilon) n \log n)=o\left(n^{2}\right)$ (when theorem applies)

[^3]
## Comparison to existing algorithms

Best existing approach: run forward or backward DP $\forall v$

- $I_{1}$ accuracy guarantee, $O\left(n^{2} \log n\right)$ complexity
- Ignores structure/dependencies across rows of $\Pi$ !

■ Our scheme accounts for structure, thus reduces complexity

Another noteworthy work: Lofgren, Banerjee, Goel 2016

- Estimates single entry of $\Pi$ via DP + MCMC, complexity $O(\sqrt{n} \log n)$
- Hence, $O\left(n^{2.5} \log n\right)$ to estimate $\Pi$; ignores dependencies across entries

■ Again, accounting for structure allows us to reduce complexity

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[^0]:    1 "Nice" = well-approximated by a certain branching process, e.g. Chen, Olvera-Cravioto 2013; Chen, Litvak, Olvera-Cravioto 2017

[^1]:    ${ }^{2}$ Can show worst-case error is $1-\alpha_{n}$, so this $\epsilon$ reduces worst-case by factor of 3

[^2]:    ${ }^{3}$ Using decomposition shown previously

[^3]:    ${ }^{4}$ Assuming $|E|=O(n), \alpha=\Theta(1 / \log n)$

