

Restart perturbations for reversible Markov chains

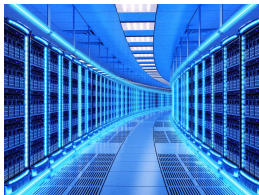
Trichotomy and pre-cutoff equivalence

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Motivation (1/2)

Discrete-time Markov chain (DTMC): a common stochastic model



Fundamental concern: sensitivity of stationary distribution to (bounded) perturbations of transition matrix

- e.g. how much do modeling errors affect long-run behavior?
- e.g. how much damage can adversary do long-term?

Motivation (2/2)

Sensitivity and mixing time

- Mixing time: number steps for DTMC to reach equilibrium/stationarity
- If fast mixing, less time for perturbation to take effect
- If slow mixing, small perturbation may have large effect

In particular, connection to cutoff

- Cutoff: far from stationary for many steps, then suddenly stationary

Table 1. Distance to stationarity for repeated shuffles of 52 cards

| | k | | | | | | | | | |
|-----------------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|
| | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| $\ P^k - \pi\ $ | 1.000 | 1.000 | 1.000 | 1.000 | 0.924 | 0.624 | 0.312 | 0.161 | 0.083 | 0.041 |

(from Diaconis 1996)

- Many examples, but little general theory

Overview

- 1 Perturbation model
- 2 Mixing times and cutoff
- 3 Illustrative examples
- 4 Results (and related work)

Basic definitions

DTMC $\{X_n(t)\}_{t \in \mathbb{Z}_+}$ with states $[n] = \{1, \dots, n\}$ and transition matrix P_n :

$$\mathbb{P}(X_n(t+1) = j | X_n(t) = i) = P_n(i, j) \quad \forall i, j \in [n], t \in \mathbb{Z}_+$$

Assume **irreducible**, **aperiodic** \Rightarrow converge to unique stationary distribution π_n from any starting state:

$$\lim_{t \rightarrow \infty} P_n^t(i, \cdot) = \pi_n = \pi_n P_n \quad \forall i \in [n]$$

Sometimes require **reversibility** (a.k.a. **local balance**):

$$\pi_n(i) P_n(i, j) = \pi_n(j) P_n(j, i) \quad \forall i, j \in [n]$$

Sometimes require **laziness** (but believe we can drop this):

$$P_n(i, i) \geq \frac{1}{2} \quad \forall i \in [n]$$

Perturbation model

Given $\alpha_n \in (0, 1)$, define

$$B(P_n, \alpha_n) = \left\{ \tilde{P}_n \in \mathcal{E}_n : \max_{i \in [n]} \|\tilde{P}_n(i, \cdot) - P_n(i, \cdot)\|_{TV} \leq \alpha_n \right\}$$

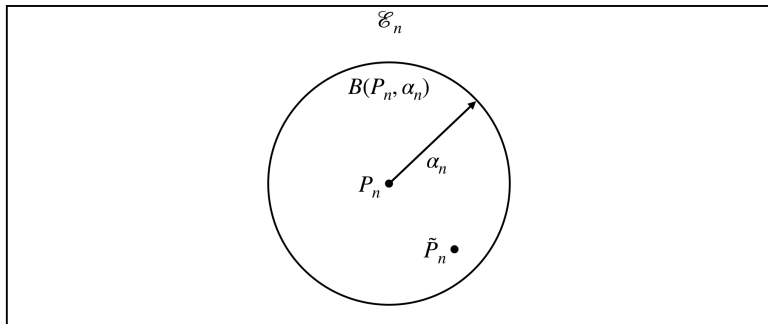
where

- $\mathcal{E}_n =$ set of $n \times n$ irreducible/aperiodic transition matrices
- $\|\mu - \nu\|_{TV} = \max_{A \subset [n]} |\mu(A) - \nu(A)|$ for distributions μ, ν on $[n]$

Denote stationary distribution of $\tilde{P}_n \in \mathcal{E}_n$ by $\tilde{\pi}_n$

Main focus of this work: relationship between $\|\tilde{\pi}_n - \pi_n\|_{TV}$ and α_n
i.e. how does steady-state error scale with perturbation magnitude?

Illustration and interpretations



| | Interpretation 1 | Interpretation 2 |
|----------------------------------|---------------------------|-------------------------------|
| P_n | True dynamics | Designed system |
| \tilde{P}_n | Modeled dynamics | System disrupted by adversary |
| α_n | Modeling accuracy | Adversary budget/ability |
| $\ \tilde{\pi}_n - \pi_n\ _{TV}$ | Effect of modeling errors | Effect of adversary |

Restart perturbations

Given $\alpha_n \in (0, 1)$ and distribution σ_n over $[n]$, define

$$P_{\alpha_n, \sigma_n} = (1 - \alpha_n)P_n + \alpha_n \mathbf{1}_n^T \sigma_n \in \mathcal{B}(P_n, \alpha_n)$$

and denote corresponding stationary distribution by π_{α_n, σ_n}

Interpretation:

- Flip Bernoulli(α_n) coin at each step
- If tails, sample next state from P_n
- If heads, sample next state from σ_n

Note: $\pi_{\alpha_n, \pi_n} = \pi_n$ (perturbations need not change stationary distribution)

Properties of restart perturbations

1 Power iteration: Solving $\pi_{\alpha_n, \sigma_n} = \pi_{\alpha_n, \sigma_n} P_{\alpha_n, \sigma_n}$ gives

$$\pi_{\alpha_n, \sigma_n} = \alpha_n \sigma_n (I_n - (1 - \alpha_n) P_n)^{-1} = \alpha_n \sigma_n \sum_{t=0}^{\infty} (1 - \alpha_n)^t P_n^t$$

2 Linearity: For distributions $\sigma_{n,1}, \sigma_{n,2}$ and $\lambda \in (0, 1)$,

$$\pi_{\alpha_n, \lambda \sigma_{n,1} + (1-\lambda) \sigma_{n,2}} = \lambda \pi_{\alpha_n, \sigma_{n,1}} + (1 - \lambda) \pi_{\alpha_n, \sigma_{n,2}}$$

3 Perfect sampling: Letting $T \sim \text{Geometric}(\alpha_n)$,

$$\pi_{\alpha_n, \sigma_n}(i) = \mathbb{P}(X_n(T) = i | X_n(0) \sim \sigma_n)$$

Sidebar: (Personalized) PageRank

When P_n is random walk on $G = ([n], E) \dots$

- $\pi_{\alpha_n, 1_n/n}$ called **PageRank**, measure of importance/centrality of nodes
- π_{α_n, e_i} called **Personalized PageRank** for i , measure of relevance to i

Applications: **Internet search** (Page et al. 1999), **recommendation systems** (Gupta et al. 2013), **community detection** (Andersen, Chung, Lang 2006; Kloumann, Ugander, Kleinberg 2017), **graph similarity** (Koutra, Vogelstein, Faloutsos 2013), **bioinformatics** (Morrison et al. 2005; Freschi 2007), ...

Rich literature: **efficient PageRank estimation** (Jeh, Widom 2003; Avrachenkov et al. 2007; Andersen et al. 2008; Lofgren, Banerjee, Goel 2016; Vial, Subramanian 2017), **implementation e.g. MapReduce** (Tong, Faloutsos, Pan 2006; Kang, Tsourakakis, Faloutsos 2009; Whang et al. 2015), **behavior on random graphs** (Avrachenkov et al. 2015; Chen, Litvak, Olvera-Cravioto 2017; Caputo, Quattropiani 2019; Vial, Subramanian 2019), ...

Mixing times

Define t -step distance from stationarity and ϵ -mixing time as

$$d_n(t) = \max_{i \in [n]} \|P_n^t(i, \cdot) - \pi_n\|_{TV}, \quad t_{\text{mix}}^{(n)}(\epsilon) = \min \{t \in \mathbb{Z}_+ : d_n(t) \leq \epsilon\}$$

Also define mixing time $t_{\text{mix}}^{(n)} = t_{\text{mix}}^{(n)}(\frac{1}{4})$

- Well-known fact: $d_n(kt_{\text{mix}}^{(n)}(\epsilon)) \leq (2\epsilon)^k$; $\epsilon = \frac{1}{4}$ gives $d_n(kt_{\text{mix}}^{(n)}) \leq 2^{-k}$

Typically desire “Big-O” behavior of $t_{\text{mix}}^{(n)}$ for sequence $\{P_n\}_{n \in \mathbb{N}}$

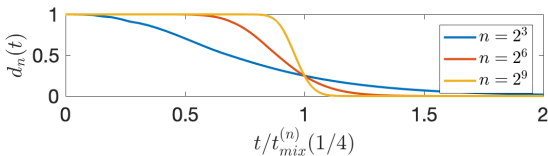
- e.g. $t_{\text{mix}}^{(n)} = \Theta(n^2)$ for random walk on n -cycle

More refined question: how “sharply” does graph of $d_n(t)$ decay to 0?

(Pre-)cutoff

Sequence of chains exhibits **cutoff** if

$$\lim_{n \rightarrow \infty} \frac{t_{\text{mix}}^{(n)}(\epsilon)}{t_{\text{mix}}^{(n)}(1 - \epsilon)} = 1 \quad \forall \epsilon \in \left(0, \frac{1}{2}\right)$$



Weaker condition of **pre-cutoff** only requires

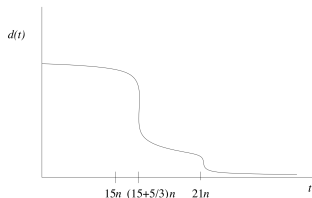
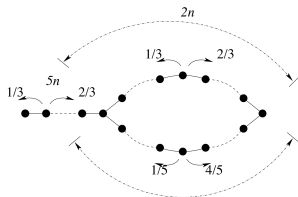
$$\sup_{\epsilon \in (0, \frac{1}{2})} \underbrace{\limsup_{n \rightarrow \infty} \frac{t_{\text{mix}}^{(n)}(\epsilon)}{t_{\text{mix}}^{(n)}(1 - \epsilon)}}_{*} < \infty$$

Literature on cutoff (1/2)

Many examples of chains with cutoff

- Classic examples (Diaconis 1996): [card shuffling](#) (e.g. Diaconis, Shahshahani 1981; Bayer, Diaconis 1992), [diffusion models](#) (e.g. Aldous 1983; Diaconis, Shahshahani 1987), ...
- Recently, cutoff for generative models, e.g. random walks on [random regular graphs](#) (Lubetzky, Sly 2010), non-backtracking walks on [sparse random graphs](#) (Ben-Hamou, Salez 2017), general models of [sparse chains](#) (Bordenave, Caputo, Salez 2019), ...

Example of pre-cutoff but no cutoff (Levin, Peres, Wilmer 2009, Ch. 18):



Literature on cutoff (2/2)

Necessary condition for pre-cutoff (and thus cutoff): inverse spectral gap vanishes relative to $t_{\text{mix}}^{(n)}$ (Levin, Peres, Wilmer 2009, Prop. 18.4)

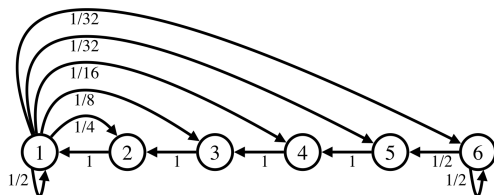
Spectral gap condition **not sufficient** (Levin, Peres, Wilmer 2009, Ex. 18.7)

No known necessary/sufficient condition in terms of eigenvalues alone (Diaconis 1996 conjectured high multiplicity of second eigenvalue)

Basu, Hermon, Peres 2015: cutoff equivalent to **"hitting time cutoff"**

Necessary/sufficient conditions for pre-cutoff?

Winning streak reversal (WSR) from Levin, Peres, Wilmer 2009



$$P_n(i, j) = \begin{cases} 2^{-j}, & i = 1, j \in \{1, \dots, n-1\} \\ 2^{-(n-1)}, & i = 1, j = n \\ 1, & i \in \{2, \dots, n-1\}, j = i-1 \\ 2^{-1}, & i = n, j \in \{n-1, n\} \end{cases}$$

$\pi_n = [2^{-1} \ 2^{-2} \ \dots \ 2^{-(n-1)} \ 2^{-(n-1)}]$, since

$$\sum_{j=1}^n \pi_n(j) P_n(j, i) = 2^{-1} \times 2^{-i} + 2^{-(i+1)} \times 1 = 2^{-i} = \pi_n(i)$$

WSR mixing behavior

Stationary by step $n - 1$, i.e. $t_{\text{mix}}^{(n)}(1 - \epsilon), t_{\text{mix}}^{(n)}(\epsilon) \leq n - 1$

- $P_n(1, \cdot) = \pi_n$ by previous slide
- For $i \in \{2, \dots, n - 1\}$, move to state 1 in $i - 1$ steps, so $P_n^i(i, \cdot) = \pi_n$
- More subtly, $P_n^{n-1}(n, \cdot) = \pi_n$

Lower bound: $t_{\text{mix}}^{(n)}(1 - \epsilon), t_{\text{mix}}^{(n)}(\epsilon) \geq n - 1 + \log_2(\epsilon)$, since

$$\begin{aligned} d_n(t) &= \max_{i \in [n], A \subset [n]} |P_n^t(i, A) - \pi_n(A)| \\ &\geq P_n^t(n - 1, \{n - 1 - t\}) - \pi_n(\{n - 1 - t\}) = 1 - 2^{-n+1+t} \end{aligned}$$

Combining gives strong form of cutoff:

$$1 \leq \frac{t_{\text{mix}}^{(n)}(\epsilon)}{t_{\text{mix}}^{(n)}(1 - \epsilon)} = 1 + O\left(\frac{1}{n}\right)$$

WSR perturbation behavior

Intuitively, “worst” perturbation is restart at n , i.e. P_{α_n, e_n}

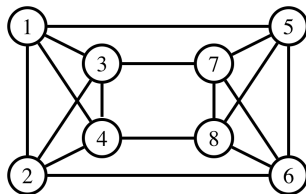
Suppose $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow \infty$ (e.g. $\alpha_n = 1/\sqrt{n}$):

- Restart at n every \sqrt{n} steps \Rightarrow rarely visit $[n/2] \Rightarrow \pi_{\alpha_n, e_n}([n/2]) \approx 0$
- Also $\pi_n([n/2]) = \sum_{i=1}^{n/2} 2^{-i} \approx 1$ (for large n)
- Taken together, $\|\pi_n - \pi_{\alpha_n, e_n}\|_{TV} \geq \pi_n([n/2]) - \pi_{\alpha_n, e_n}([n/2]) \approx 1$

If instead $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow 0$ (e.g. $\alpha_n = 1/n^2$):

- Visit state 1 before restart, reach original stationary distribution
- Once “locked into” this distribution, can’t escape
- Consequently, $\|\pi_n - \pi_{\alpha_n, e_n}\|_{TV} \approx 0$

Complete graph bijection (CGB)



Construction:

- Draw complete graphs / cliques on $\{1, \dots, \frac{n}{2}\}$ and $\{1 + \frac{n}{2}, \dots, n\}$
- Add edges $\{(1, 1 + \frac{n}{2}), (2, 2 + \frac{n}{2}), \dots, (\frac{n}{2}, n)\}$

We study lazy simple random walk on resulting graph

By symmetry, uniform distribution is stationary

CGB mixing behavior

“Half-mixed” in one step, i.e. $t_{\text{mix}}^{(n)}(1 - \epsilon) = \Theta(1)$

- $P_n(X_n(0), \cdot) \approx$ uniform on $X_n(0)$'s clique $\Rightarrow \|P_n(X_n(0), \cdot) - \pi_n\|_{TV} \approx \frac{1}{2}$

To fully mix, require n steps, i.e. $t_{\text{mix}}^{(n)}(\epsilon) = \Theta(n)$

- At each step, $\Theta(\frac{1}{n})$ mass escapes $X_n(0)$'s clique
- Thus, $\Theta(n)$ steps for constant mass to escape $X_n(0)$'s clique

“Opposite” of cutoff, since

$$\frac{t_{\text{mix}}^{(n)}(\epsilon)}{t_{\text{mix}}^{(n)}(1 - \epsilon)} = \Theta(n)$$

which is maximal (up to constants) among $t_{\text{mix}}^{(n)}(\epsilon) = \Theta(n)$ chains

CGB perturbation behavior

By symmetry, “worst” (restart) perturbation is P_{α_n, e_1}

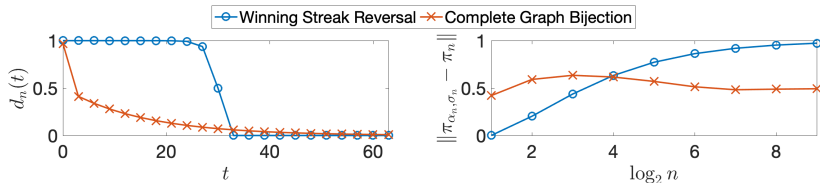
One step after restart, uniform on $\{1, \dots, \frac{n}{2}\}$

“Half-mixed”, can’t fully escape this distribution

More specifically, can show that for any $\alpha_n \rightarrow 0$ and any $\tilde{P}_n \in B(P_n, \alpha_n)$,

$$\limsup_{n \rightarrow \infty} \|\tilde{\pi}_n - \pi_n\|_{TV} \leq \frac{1}{2}$$

Summary and key insights



WSR: strong form of cutoff, sensitive to α_n -magnitude perturbations

- Assuming $\alpha_n \gg 1/t_{\text{mix}}^{(n)}(\epsilon)$

CGB: “opposite” of cutoff, robust against perturbation

Do these phenomena hold more generally?

Trichotomy result

Theorem (Vial, Subramanian 2019)

Let $P_n \in \mathcal{E}_n$, $\alpha_n \in (0, 1)$, $\epsilon \in (0, 1)$. Assume $\lim_{n \rightarrow \infty} \alpha_n t_{mix}^{(n)}(\epsilon) = c \in [0, \infty]$, P_n is lazy and reversible $\forall n$, and $\{P_n\}$ has cutoff.

- If $c = 0$, then $\forall \{\tilde{P}_n\}$ with $\tilde{P}_n \in B(P_n, \alpha_n) \forall n$,

$$\lim_{n \rightarrow \infty} \|\tilde{\pi}_n - \pi_n\|_{TV} = 0.$$

- If $c = \infty$, then $\exists \{\sigma_n\}$ with $\sigma_n \in \Delta_{n-1} \forall n$ s.t.

$$\lim_{n \rightarrow \infty} \|\pi_{\alpha_n, \sigma_n} - \pi_n\|_{TV} = 1.$$

- If $c \in (0, \infty)$, then $\forall \{\sigma_n\}$ with $\sigma_n \in \Delta_{n-1} \forall n$,

$$\limsup_{n \rightarrow \infty} \|\pi_{\alpha_n, \sigma_n} - \pi_n\|_{TV} \leq 1 - e^{-c}.$$

Also, the bound is tight, i.e. $\exists \{\sigma_n\}$ with $\sigma_n \in \Delta_{n-1} \forall n$ s.t.

$$\liminf_{n \rightarrow \infty} \|\pi_{\alpha_n, \sigma_n} - \pi_n\|_{TV} \geq 1 - e^{-c}.$$

Notes on theorem

Tight characterization: given any $c' \in [0, 1]$, can choose α_n, σ_n s.t.

$$\lim_{n \rightarrow \infty} \|\pi_{\alpha_n, \sigma_n} - \pi_n\|_{TV} = c'.$$

Some bounds require weaker assumptions (see Lemmas 1-2 in paper)

- Upper bounds don't need laziness, reversibility, or cutoff
- Lower bounds only need pre-cutoff in the case $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow \infty$

Notes on assumptions for lower bound:

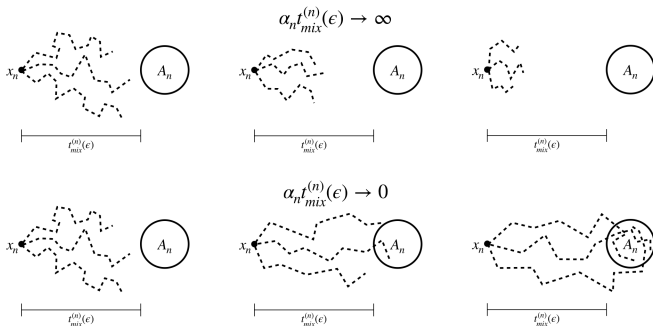
- Believe we can drop laziness
- Reversibility common in mixing times literature (real eigenvalues)
- Necessity of (pre-)cutoff – next theorem

Proof idea

Basu, Hermon, Peres 2015: if pre-cutoff holds, $\exists x_n \in [n], A_n \subset [n]$ s.t.

$$\pi_n(A_n) \approx 1, \quad t_{\text{mix}}^{(n)}(\epsilon) \propto \text{time to reach } A_n \text{ from } x_n$$

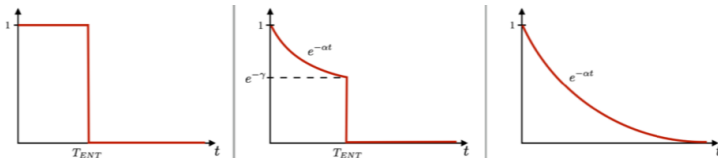
i.e. “mixing times are hitting times of large sets” (also Peres, Sousi 2015)



Related results in literature (1/2)

Caputo, Quattropani 2019 shows distance to stationarity $d_{\alpha_n, \sigma_n}(t)$ on P_{α_n, σ_n} follows trichotomy when $P_n =$ random walk on configuration model (CM):

- If $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow 0$, d_{α_n, σ_n} is step function, like d_n (left)
- If $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow \infty$, $d_{\alpha_n, \sigma_n}(t)$ decays exponentially (right)
- If $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow (0, \infty)$, interpolation (middle)



Avena et al. 2018: similar result for dynamic CM

Related results in literature (2/2)

Vial, Subramanian 2019 studies matrix $\Pi = \{\pi_{\alpha_n, e_i}\}_{i \in [n]}$ for CM

- If $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow 0$, $\text{rank}(\Pi) = \Theta(1)$
- If $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow \infty$, $\text{rank}(\Pi) = \Omega(\frac{n}{\log n})$ (conjecture)
- If $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow (0, \infty)$, $\text{rank}(\Pi) = O(n^c)$, $c \in (0, 1)$

These four results have similar flavor:

- If $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow 0$, some property unchanged
- If $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow \infty$, this property maximally changed
- If $\alpha_n t_{\text{mix}}^{(n)}(\epsilon) \rightarrow (0, \infty)$, intermediate behavior

Suggests (ill-posed) question: characterizing this class of properties

Pre-cutoff “equivalence”

Some notion of cutoff assumed for all of our lower bounds – is it necessary?

Sensitivity Condition (SC)

For *certain sequences* $\{\alpha_{n,\epsilon}\}_{n \in \mathbb{N}, \epsilon \in (0,1/2)} \subset (0,1)$, $\exists \{\sigma_{n,\epsilon}\}_{n \in \mathbb{N}, \epsilon \in (0,1/2)}$ such that $\lim_{n \rightarrow \infty} \|\pi_n - \pi_{\alpha_{n,\epsilon}, \sigma_{n,\epsilon}}\|_{TV} = 1 \forall \epsilon \in (0,1/2)$.

Theorem (Vial, Subramanian 2019)

Assume each P_n is lazy and reversible $\forall n$. If pre-cutoff holds, then SC holds. If

$$\sup_{\epsilon \in (0,1/2)} \liminf_{n \rightarrow \infty} t_{mix}^{(n)}(\epsilon) / t_{mix}^{(n)}(1 - \epsilon) = \infty, \quad (1)$$

then SC fails.

- (1) slightly stronger than converse of pre-cutoff (lim inf instead of lim sup)
- Thus, theorem says **pre-cutoff and SC (almost) equivalent**
- Complements Basu, Hermon, Peres 2015 (cutoff \Leftrightarrow “hitting time cutoff”)

Illustration of “equivalence”

Partition of lazy/reversible sequences induced by SC:

| SC holds | SC fails |
|--|--|
| $\sup_{\epsilon \in (0, 1/2)} \limsup_{n \rightarrow \infty} \frac{t_{mix}^{(n)}(\epsilon)}{t_{mix}^{(n)}(1 - \epsilon)} < \infty$ | $\sup_{\epsilon \in (0, 1/2)} \liminf_{n \rightarrow \infty} \frac{t_{mix}^{(n)}(\epsilon)}{t_{mix}^{(n)}(1 - \epsilon)} = \infty$ |

Believe gray area only contains “bizarre” sequences

- e.g. alternating between WSR and CGB

“Certain sequences”

Sensitivity Condition (SC)

For any $\{\alpha_{n,\epsilon}\}_{n \in \mathbb{N}, \epsilon \in (0,1/2)} \subset (0,1)$ satisfying (2), $\exists \{\sigma_{n,\epsilon}\}_{n \in \mathbb{N}, \epsilon \in (0,1/2)}$ such that $\lim_{n \rightarrow \infty} \|\pi_n - \pi_{\alpha_{n,\epsilon}, \sigma_{n,\epsilon}}\|_{TV} = 1 \forall \epsilon \in (0,1/2)$.

$$\sup_{\epsilon \in (0,1/2)} \liminf_{n \rightarrow \infty} \alpha_{n,\epsilon} t_{mix}^{(n)}(\epsilon) = \infty, \frac{\alpha_{n,\epsilon}}{\alpha_{n,\delta}} \in \left[\frac{t_{mix}^{(n)}(1-\delta)}{t_{mix}^{(n)}(1-\epsilon)}, 1 \right] \forall 0 < \delta \leq \epsilon < \frac{1}{2}. \quad (2)$$

Proof intuition broadly similar to previous theorem

Little intuition for (2), but yields the following “magic”:

- When pre-cutoff **fails**, $\{\{\alpha_{n,\epsilon}\}_{n,\epsilon} : (2) \text{ holds}\}$ is **sufficiently large** that we can find $\{\sigma_{n,\epsilon}\}_{n,\epsilon}$ **violating** conclusion of SC
- When pre-cutoff **holds**, $\{\{\alpha_{n,\epsilon}\}_{n,\epsilon} : (2) \text{ holds}\}$ **collapses** to a set for which conclusion of SC **holds**

Conclusions

Main takeaways

- Perturbation error depends on perturbation magnitude \times mixing time
- Pre-cutoff almost equivalent to certain notion of “perturbation sensitivity”

Further reading

- Vial, Subramanian, “Restart Perturbations for Lazy, Reversible Markov Chains: Trichotomy and Pre-cutoff Equivalence”, arXiv 1907.02926
- Vial, Subramanian, “A Structural Result for Personalized PageRank and its Algorithmic Consequences”, POMACS, June 2019
- Levin, Peres, Wilmer, “Markov Chains and Mixing Times”
- Diaconis, “The Cutoff Phenomenon in Finite Markov Chains”, PNAS February 1996
- Basu, Hermon, Peres, “Characterization of cutoff for reversible Markov chains”, SODA 2015

- Aldous, David (1983). "Random walks on finite groups and rapidly mixing Markov chains". In: *Séminaire de Probabilités XVII 1981/82*. Springer, pp. 243–297.
- Andersen, Reid, Fan Chung, Kevin Lang (2006). "Local graph partitioning using PageRank vectors". In: *2006 47th Annual IEEE Symposium on Foundations of Computer Science (FOCS'06)*. IEEE, pp. 475–486.
- Andersen, Reid et al. (2008). "Local computation of PageRank contributions". In: *Internet Mathematics* 5.1-2, pp. 23–45.
- Avena, Luca et al. (2018). "Random walks on dynamic configuration models: a trichotomy". In: *Stochastic Processes and their Applications*.
- Avrachenkov, Konstantin et al. (2007). "Monte Carlo methods in PageRank computation: When one iteration is sufficient". In: *SIAM Journal on Numerical Analysis* 45.2, pp. 890–904.
- Avrachenkov, Konstantin et al. (2015). "PageRank in undirected random graphs". In: *International Workshop on Algorithms and Models for the Web-Graph*. Springer, pp. 151–163.
- Basu, Riddhipratim, Jonathan Hermon, Yuval Peres (2015). "Characterization of cutoff for reversible Markov chains". In: *Proceedings of the twenty-sixth annual ACM-SIAM symposium on Discrete algorithms*. Society for Industrial and Applied Mathematics, pp. 1774–1791.
- Bayer, Dave, Persi Diaconis, et al. (1992). "Trailing the dovetail shuffle to its lair". In: *The Annals of Applied Probability* 2.2, pp. 294–313.
- Ben-Hamou, Anna, Justin Salez, et al. (2017). "Cutoff for nonbacktracking random walks on sparse random graphs". In: *The Annals of Probability* 45.3, pp. 1752–1770.

- Bordenave, Charles, Pietro Caputo, Justin Salez (2019). “Cutoff at the “entropic time” for sparse Markov chains”. In: *Probability Theory and Related Fields* 173.1-2, pp. 261–292.
- Caputo, Pietro, Matteo Quattropiani (2019). “Mixing time of PageRank surfers on sparse random digraphs”. In: *arXiv preprint arXiv:1905.04993*.
- Chen, Ningyuan, Nelly Litvak, Mariana Olvera-Cravioto (2017). “Generalized PageRank on directed configuration networks”. In: *Random Structures & Algorithms* 51.2, pp. 237–274.
- Diaconis, Persi (1996). “The cutoff phenomenon in finite Markov chains”. In: *Proceedings of the National Academy of Sciences* 93.4, pp. 1659–1664.
- Diaconis, Persi, Mehrdad Shahshahani (1981). “Generating a random permutation with random transpositions”. In: *Probability Theory and Related Fields* 57.2, pp. 159–179.
- (1987). “Time to reach stationarity in the Bernoulli–Laplace diffusion model”. In: *SIAM Journal on Mathematical Analysis* 18.1, pp. 208–218.
- Freschi, Valerio (2007). “Protein function prediction from interaction networks using a random walk ranking algorithm”. In: *Bioinformatics and Bioengineering, 2007. BIBE 2007. Proceedings of the 7th IEEE International Conference on*. IEEE, pp. 42–48.
- Gupta, Pankaj et al. (2013). “WTF: The who to follow service at twitter”. In: *Proceedings of the 22nd international conference on World Wide Web*. ACM, pp. 505–514.
- Jeh, Glen, Jennifer Widom (2003). “Scaling personalized web search”. In: *Proceedings of the 12th international conference on World Wide Web*. ACM, pp. 271–279.

- Kang, U, Charalampos E Tsourakakis, Christos Faloutsos (2009). "Pegasus: A peta-scale graph mining system implementation and observations". In: *Proceedings of the 2009 Ninth IEEE International Conference on Data Mining*. Washington, DC, USA, pp. 229–238.
- Kloumann, Isabel M, Johan Ugander, Jon Kleinberg (2017). "Block models and personalized PageRank". In: *Proceedings of the National Academy of Sciences* 114.1, pp. 33–38.
- Koutra, Danai, Joshua T Vogelstein, Christos Faloutsos (2013). "Deltacon: A principled massive-graph similarity function". In: *Proceedings of the 2013 SIAM International Conference on Data Mining*. SIAM, pp. 162–170.
- Levin, David A, Yuval Peres, Elizabeth L Wilmer (2009). *Markov chains and mixing times*. American Mathematical Society.
- Lofgren, Peter, Siddhartha Banerjee, Ashish Goel (2016). "Personalized PageRank estimation and search: A bidirectional approach". In: *Proceedings of the Ninth ACM International Conference on Web Search and Data Mining*. ACM, pp. 163–172.
- Lubetzky, Eyal, Allan Sly, et al. (2010). "Cutoff phenomena for random walks on random regular graphs". In: *Duke Mathematical Journal* 153.3, pp. 475–510.
- Morrison, Julie L et al. (2005). "GeneRank: Using search engine technology for the analysis of microarray experiments". In: *BMC bioinformatics* 6.1, p. 233.
- Page, Lawrence et al. (1999). *The PageRank citation ranking: Bringing order to the web*. Tech. rep. Stanford InfoLab.
- Peres, Yuval, Perla Sousi (2015). "Mixing times are hitting times of large sets". In: *Journal of Theoretical Probability* 28.2, pp. 488–519.

- Tong, Hanghang, Christos Faloutsos, Jia-Yu Pan (2006). “Fast random walk with restart and its applications”. In: *Sixth International Conference on Data Mining (ICDM'06)*. IEEE, pp. 613–622.
- Vial, Daniel, Vijay Subramanian. “Restart perturbations for lazy, reversible Markov chains: trichotomy and pre-cutoff equivalence”. In: *(In Preparation)* ().
- (2017). “On the role of clustering in Personalized PageRank estimation”. In: *arXiv preprint arXiv:1706.01091*.
 - (2019). “A Structural Result for Personalized PageRank and its Algorithmic Consequences”. In: *Proceedings of the ACM on Measurement and Analysis of Computing Systems* 3.2, p. 25.
- Whang, Joyce Jiyoung et al. (2015). “Scalable data-driven pagerank: Algorithms, system issues, and lessons learned”. In: *European Conference on Parallel Processing*. Springer, pp. 438–450.