

# Econ 679 – Midterm Exam

MATIAS D. CATTANEO  
DEPARTMENT OF ECONOMICS  
UNIVERSITY OF MICHIGAN

October 24, 2008

## CONTENTS

1	Problem 1 – Asymmetric Least Squares . . . . .	1
2	Problem 2 – Efficient Two-step Estimation . . . . .	2

### 1. PROBLEM 1 – ASYMMETRIC LEAST SQUARES

Consider the following (robust) least squares based alternative to quantile regression:

$$\hat{\beta}_\pi = \arg \min_{\beta} M_n(\beta), \quad M_n(\beta) = \frac{1}{n} \sum_{i=1}^n (Y_i - X_i' \beta)^2 |\pi - \mathbf{1}(Y_i < X_i' \beta)|,$$

where  $\pi \in (0, 1)$  and  $(Y_i, X_i)'$ ,  $i = 1, 2, \dots, n$ , are i.i.d. satisfying the linear regression relationship  $Y_i = X_i' \beta_\pi^* + \varepsilon_i$ , where  $\varepsilon_i$  are independent of  $X_i$  for all  $i = 1, 2, \dots, n$ .

1. Assuming that  $X_i$  contains an intercept and the normalization

$$0 = \arg \min_{m \in \mathbb{R}} \mathbb{E} \left[ (\varepsilon_i - m)^2 |\pi - \mathbf{1}(\varepsilon_i < m)| \right]$$

holds, provide sufficient conditions for identifiability of  $\beta_\pi^*$  (as the solution to the limiting M-estimation problem).

2. Provide sufficient conditions so that  $\hat{\beta}_\pi = \beta_\pi^* + o_p(1)$ .
3. Provide sufficient conditions so that

$$\sqrt{n}(\hat{\beta}_\pi - \beta_\pi^*) = \frac{1}{\sqrt{n}} \sum_{i=1}^n \psi(Y_i, X_i; \beta_\pi^*) + o_p(1),$$

and conclude that  $\sqrt{n}(\hat{\beta}_\pi - \beta_\pi^*) \xrightarrow{d} \mathcal{N}(0, V_\pi)$ . Identify the exact form of the influence function  $\psi(Y_i, X_i; \beta_\pi^*)$  and limiting covariance matrix  $V_\pi$ .

4. Note that  $\pi = 1/2$  corresponds to the classical least squares (LS) regression estimator. Thus, verify that the limiting distribution derived in part 3 agrees with the limiting distribution of LS if  $\pi = 1/2$ .

## 2. PROBLEM 2 – EFFICIENT TWO-STEP ESTIMATION

Suppose that the (scalar) parameter of interest  $\beta_0$  solves the moment condition

$$\mathbb{E}[m_1(X, \beta_0, \gamma_0)] = 0,$$

where  $\gamma_0$  is some other unknown (scalar) parameter, which solves the moment condition

$$\mathbb{E}[m_2(X, \gamma_0)] = 0.$$

Assume that  $m(X, \beta, \gamma)$  is continuously differentiable as needed, and a random sample of observations  $X_1, \dots, X_n$  is available.

1. Suppose  $\gamma_0$  is known and, using this information, consider the estimator  $\tilde{\beta}_n$  given by

$$\frac{1}{n} \sum_{i=1}^n m_1(X_i, \tilde{\beta}_n, \gamma_0) = 0.$$

Provide sufficient conditions for consistency and derive the limiting distribution of  $\tilde{\beta}_n$ .

2. Suppose  $\gamma_0$  is unknown and consider the two-step estimator  $\hat{\beta}_n$  given by

$$\frac{1}{n} \sum_{i=1}^n m_1(X_i, \hat{\beta}_n, \hat{\gamma}_n) = 0,$$

where  $\hat{\gamma}_n$  is given by

$$\frac{1}{n} \sum_{i=1}^n m_2(X_i, \hat{\gamma}_n) = 0.$$

Provide sufficient conditions for consistency and derive the (joint) limiting distribution of  $(\hat{\beta}_n, \hat{\gamma}_n)'$ .

3. Assume that  $\mathbb{E}[m_1(X, \beta_0, \gamma_0) m_2(X, \gamma_0)] = 0$ . Using the previous two results, verify that  $\mathbb{V}[\hat{\beta}_n] \geq \mathbb{V}[\tilde{\beta}_n]$  and derive (alternative) conditions so that  $\mathbb{V}[\hat{\beta}_n] = \mathbb{V}[\tilde{\beta}_n]$  (i.e., the feasible estimator is as efficient as the unfeasible one).
4. Assume that  $\mathbb{E}[m_1(X, \beta_0, \gamma_0) m_2(X, \gamma_0)] \neq 0$ . Is the unfeasible estimator (i.e., the one that uses the information about  $\gamma_0$ ) always as efficient as  $\hat{\beta}_n$ ? Provide intuition for your answer.
5. In conclusion, assuming  $\gamma_0$  is known, how would you estimate  $\beta_0$  efficiently? Is this result intuitive?